



Stabilizing Error Correction Mechanism in the Presence of Explosiveness

Modupe O. Abifarin ^{a*}, Audu Isah ^a, Yisa Yakubu ^a
and Rasheed A. Adeyemi ^a

^a Department of Statistics, Federal University of Technology, Minna, Niger State, Nigeria.

Authors' contributions

This work was carried out in collaboration among all authors. Author AMO designed the study, performed the statistical analysis, wrote the protocol and wrote the first draft of the manuscript, Authors IA, YY and ARA modified the derivatives and corrected the draft of the manuscript. All authors read and approved the final manuscript.

Article Information

DOI: 10.9734/AJPAS/2024/v26i1584

Open Peer Review History:

This journal follows the Advanced Open Peer Review policy. Identity of the Reviewers, Editor(s) and additional Reviewers, peer review comments, different versions of the manuscript, comments of the editors, etc are available here: <https://www.sdiarticle5.com/review-history/112368>

Original Research Article

Received: 19/11/2023

Accepted: 25/01/2024

Published: 01/02/2024

Abstract

In the presence of explosiveness of the adjustment term in the error correction model, the adjustment of the dependent variable Y was too large and overshoots the equilibrium, creating a divergent pattern. The error correction model fails to capture the deviation from equilibrium appropriately, thereby resulting in overshooting of the model. In this paper, a new model to stabilize the explosiveness in an Error Correction model called the stabilizing Error Correction Mechanism was proposed. Mathematical methodology for obtaining the estimate of the model using the Ordinal Least Square method was derived. Error Correction model was used to model the relationship among the variables and the result was compared with the Stabilizing Error Correction Mechanism using root mean square error. A Monte-Carlo simulation was performed, and the stimulation results showed that the error correction model exhibited some explosiveness, and the damping coefficient of the stabilizing model exerted a stabilizing effect on the error correction mechanism, thereby reducing the overshooting in the error correction model. The proposed model contributed to a smoother and more stable response to deviations from the long-run equilibrium. The root mean square

*Corresponding author: Email: olaniretimodupe@gmail.com;

error of the stabilizing Error Correction model was observed to be 1.30663, 1.04533, 12.55786, 10.49876, 10.0034, and 19.41545 as compared to the adjustment model in the Error Correction model (60.6888, 35.5929, 315238, 24.31958, 10.1485 and 19.7687) when the persistence is high and $(\beta_0, \beta_1, \phi) = (1, 0, 0)$. Therefore, the Stabilizing Error Correction model performs better than the Error Correction model.

Keywords: Autoregressive distributed lag; error correction model; long-run model and stabilizing error correction model.

1 Introduction

Econometrics analysis of long-run relationship has been the focus of much theoretical and empirical research in economics, Pesaran and Shin [1]. The Autoregressive Distributed lag (ARDL) model is adopted when the dynamics of a single equation regression is involved, when the variables are non-stationary, the ARDL model is reparameterized into an error correction form to study the short-run fluctuation around the equilibrium. Attempt to uncover the long-run/equilibrium relationship is equivalent to separating it from its short-term dynamics which shows evidence for/against the equilibrium relationship Kripfganz and Schneider [2]. The ARDL model can be applied to study the relationship between different economic indicators such as gross domestic product, exchange rate, money supply, inflation, and interest rate over time, Pesaran et al. [3], Adamu and Usman [4], Aronu et al. [5], Ibrahim [6], Charles et al. [7], Elem-Uche et al. [8]; to examine the relationship between financial variables such as stock prices and economic indicators, Adeleye et al. [9], Narayan and Smyth [10], Catau and Asmah [11], Mustafa [12], Celina, U.C. [13], Enisan and Olufisayo [14], Rostin et al. [15], Liaqat et al. [16]; to investigate the impact of exchange rate on trade balance; Bahmani and Narayan [17], Belloumi [18]; to examine the effect of environmental factors, policies, or regulations on economic variables, Ozturk and Acaravci [19], Hamid et al. [20], Saida and Kais [21]; to study the long-run and short-term effect of healthcare policies, expenditures, and other health-related variables, Mamun and Sohag [22]; to examine the relationship between energy prices, consumption, and economic growth over time, Zhigang and Huang [23]. The idea of differencing of integrated time series before modelization was advocated by Box and Jenkins [24] while Engle and Granger [25] formalized the idea of cointegration, which is used in a variety of economic models (Iyeli et al., [26], Nkoro and Uko [27]). In recent years, the cointegration method has been developed to address the issue of non-stationarity in time-series data. Over the past forty years, it has established itself as a standard tool in econometrics and provides evidence for the presence of a real long-term economic relationship. Applying the real economic data to the cointegration test gives a formal and practical foundation for evaluating the short-run and long-term models. When two or more economic variables are cointegrated, short-term deviations from equilibrium have an effect on the other variables, which in turn influences a shift in the direction of the long-run equilibrium. The cointegration of the two variables suggests that an adjustment mechanism is in place to keep the long-run relationship's errors from increasing. The Error Correction model is used to measure the correction from the disequilibrium of the previous period, which has a very good economic implication; it eliminates trends from the variables involved and thereby resolves the problem of spurious regressions. Muritala et al. [28], Abdulaziz and Basmah [29], and Wasanthi [30] observed some explosiveness in the adjustment term creating a divergent pattern where the error correction model fails to capture the deviation from equilibrium appropriately, thereby resulting in overshooting of the model.

The objective of this paper is to investigate the explosiveness in the Error correction model and to propose a model to adjust for the explosiveness in the model. The paper is limited to the short-run error correction model. The plan of the paper is as follows: Section 2 provides the methods to be applied, which include the ARDL, cointegration, and the Error Correction Mechanism. Section 3 adjusts for explosiveness in the model. Section 4 presents the Monte-Carlo simulation results for the model, and Section 5 is on some concluding remarks

2 Materials and Methods

2.1 The Autoregressive Distributed Lag (ARDL) model

The Autoregressive Distributed Lag (ARDL) model by Pesaran and Shin [1] is given as:

$$\alpha(L)Y_t = \mu + \beta(L)X_t + \mu_t \quad (1)$$

Where,

X_t is explanatory variable, Y_t is the dependent variable, and μ_t is the stationary error term,
 $\beta(L)$ and $\alpha(L)$ are the lag polynomials such that;
 $\alpha(L) = 1 - \alpha_1 L - \alpha_2 L^2 - \dots - \alpha_p L^p$
 $\beta(L) = 1 - \beta_1 L - \beta_2 L^2 - \dots - \beta_q L^q$
 We assume that X_t is vector of a random variable and
 $X_t = X_{t-1} + e_t$ (2)

Where X_{t-1} is the lag of X_t , e_t is stationary and equation (2.2) implies that X_t is integrated of $I(1)$ or $I(0)$ and Y_t is also of Order $I(1)$, Y_t and X_t are cointegrated, allowing for the $Cov(u_t, e_t) \neq 0$ in which X_t is said to be endogenous

2.2 Cointegration model

Equation (1) can be rewritten as

$$\alpha(1)Y_t = \mu + [\beta(1) + \beta'(L) - \beta(1)]X_t + [\alpha(1) - \alpha(L)]Y_t + \mu_t$$

$$Y_t = \frac{\mu}{\alpha(1)} + \frac{\beta(1)'}{\alpha(1)}X_t + \frac{[\beta(L) - \beta(1)]'(1-L)}{\alpha(1)(1-L)}X_t + \frac{[\alpha(1) - \alpha(L)](1-L)}{\alpha(1)(1-L)}Y_t + \frac{\mu_t}{\alpha(1)}$$

$$Y_t = \lambda_1 + \lambda_2'X_t + \gamma_2'(L)\Delta X_t + \gamma_1(L)\Delta Y_t + v_t \quad (3)$$

Where

$$\lambda_1 = \frac{\mu}{\alpha(1)}, \lambda_2 = \frac{\beta(1)'}{\alpha(1)}$$

$$\gamma_2' = \frac{[\beta(L) - \beta(1)]'}{\alpha(1)(1-L)}, \gamma_1 = \frac{[\alpha(1) - \alpha(L)]}{\alpha(1)(1-L)} \text{ and } v_t = \frac{\mu_t}{\alpha(1)}$$

Equation (3) is the cointegration model and λ_2 measures the long-run impact of X on Y .

Stock [31], and Engle and Granger [25] estimated equation (3) using OLS estimator since Y_t and X_t are of order $I(1)$ and ΔY_t and ΔX_t are of order $I(0)$.

2.3 Error correction model

Using (ARDL (1,1)), we have:

$$Y_t = \alpha_0 + \alpha_1 Y_{t-1} + \beta_0 x_t + \beta_1 x_{t-1} + \varepsilon_t \quad (4)$$

$$Y_t - \alpha_1 Y_{t-1} = \alpha_0 + \beta_0 x_t + \beta_1 x_{t-1} + \varepsilon_t \quad (5)$$

$$Y_t(1 - \alpha_1) = \alpha_0 + (\beta_0 + \beta_1)x_t + \varepsilon_t \quad (6)$$

$$Y_t = \frac{\alpha_0}{(1-\alpha_1)} + \frac{(\beta_0 + \beta_1)x_t}{(1-\alpha_1)} + \frac{\varepsilon_t}{(1-\alpha_1)} \quad (7)$$

$$Y_t = \alpha_0^* + \beta^* x_t + \varepsilon_t \quad (8)$$

Equation (8) is the long-run multiplier.

Re-parameterizing the ARDL model using the technique of Perasan and Shin [32], is done by adding and subtracting Y_{t-1} and $\beta_0 x_{t-1}$ into equation 4, we have

$$Y_t = \alpha_0 + \alpha_1 Y_{t-1} + Y_{t-1} - Y_{t-1} + \beta_0 x_t + \beta_0 x_{t-1} - \beta_0 x_{t-1} + \beta_1 x_{t-1} + \varepsilon_t \quad (9)$$

Rearranging we have,

$$Y_t - Y_{t-1} = \alpha_o + \alpha_1 Y_{t-1} - Y_{t-1} + \beta_0 x_t - \beta_0 x_{t-1} + \beta_0 x_{t-1} + \beta_1 x_{t-1} + \varepsilon_t \tag{10}$$

$$\Delta Y_t = \alpha_o + Y_{t-1}(\alpha_1 - 1) + \beta_0 \Delta x_t + (\beta_0 + \beta_1)x_{t-1} + \varepsilon_t \tag{11}$$

$$\Delta Y_t = -(1 - \alpha_1) \left[Y_{t-1} - \frac{\alpha_o}{(1-\alpha_1)} - \frac{(\beta_0+\beta_1)}{(1-\alpha_1)} x_{t-1} \right] + \beta_0 \Delta x_t + \varepsilon_t \tag{12}$$

The interpretation of the error correction model relies on a long run equilibrium relationship $y = \beta'x$. The error correction mechanism is the adjustment of y_t through a(1) to equilibrium deviation in the previous period, in equation (8). The equation is rewritten as follows.

$$\Delta y_t = \lambda ECM_{t-1} + \beta_0 \Delta x_t + \varepsilon_t \tag{13}$$

Where $ECM_{t-1} = \left[Y_{t-1} - \frac{\alpha_o}{(1-\alpha_1)} - \frac{(\beta_0+\beta_1)}{(1-\alpha_1)} x_{t-1} \right]$ and $\lambda = -(1 - \alpha_1)$

Where $-1 < \lambda < 0$

$$\lambda = \left(\begin{array}{l} > 0, \text{ the disequilibrium expands} \\ = 0, \text{ there is no error correction} \\ -1 < \lambda < 0, \text{ a quick equilibrium} \\ = -1, \text{ Full error correction in one point} \\ < -1, \text{ over shooting; oscillatory adjustment} \end{array} \right)$$

3 Explosiveness in the Error Correction Model

In the figure below, the blue line represents the long-run equilibrium relationship between Y and X, and the red line represents the actual values of Y. The green horizontal line show the direction of adjustment of Y in each period. When λ is the coefficient of the error correction term (ECT) which is expected to be between -1 and 0. The negative indicate the degree of correction. The magnitude of the correction suggests a fairly high speed of adjustment in the aftermath of a shock. When λ is less than -1, the adjustment of Y is too large and overshoots the equilibrium, creating a divergent pattern. This is in contrast to the case when λ is between -1 and 0, where the adjustment of Y is gradual and convergent, and the red line eventually approaches the blue line. To correct the problem of explosiveness in the short run model a stabilizing error correction mechanism is proposed which is given as:

Stabilizing Error Correction Mechanism.

$$\Delta Y_t = \sum_{i=1}^p \beta_i \Delta Y_{t-i} + \sum_{j=0}^q \beta_j \Delta X_{t-j} + \lambda ECT_{t-1} + \theta SECM + e_t \tag{14}$$

SECM is called the Stabilization error correction mechanism by damping the speed of correction. It measures the change in the error correction term from one period to the next. It helps to slow down the correction process. By introducing SECM, we allow the model to adjust the speed of correction dynamically based on past adjustments.

In Matrix form equation (14) can be represented as:

$$y_t = h d_t + \theta SECM_t + e_t \tag{15}$$

$$y_t = [y_1, y_2, \dots, y_n]'$$

$$y_{t-1} = [y_0, y_1, \dots, y_{n-1}]'$$

$$\begin{aligned} e_t &= [e_1, e_2, \dots, e_n] \\ d_t &= [\Delta Y_t \ \Delta X_t \ ECM_t] \\ g &= [\beta, \theta] \end{aligned}$$

set

$$\begin{aligned} H_{dt} &= [d_t, SECM_t] \\ G &= [g, \lambda] \end{aligned}$$

Using OLS, the estimator G is denoted by G_t

$$G_t = [H_{dt}' H_{dt}]^{-1} [H_{dt}' y_t] \tag{16}$$

$$V(d) = \sigma_u^2 [H_{dt}' H_{dt}]^{-1} \tag{17}$$

$$\text{And } \sigma_u^2 = N^{-1} (y - H_{dt}' \widehat{G}_t)' (y - H_{dt}' \widehat{G}_t) \tag{18}$$

The t- statistic for the parameter estimate is given as

$$t = \frac{G_t - \widehat{G}_t}{S.E(\widehat{G}_t)} \tag{19}$$

$$S.E(\widehat{G}_t) = \sqrt{\text{var}(G_t)} \tag{20}$$

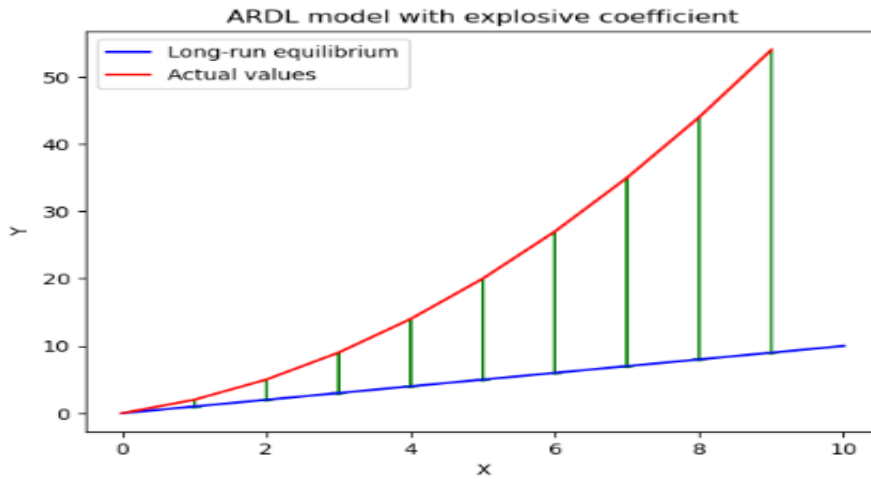


Fig. 1. Illustration of explosiveness

4 Monte-Carlo Simulation

In this section, Monte-Carlo analysis used to illustrate the impact of the shock. The sample size considered is 50.

4.1 Data generating process

The parameters are selected arbitrarily.

$$y_t = \alpha + \beta_0 x_t + \beta_1 x_{t-1} + \phi y_{t-1} + u_t \tag{21}$$

Where,

$$x_t = \rho_1 x_{t-1} + v_t \tag{22}$$

$\eta_{1t} \sim N(0,1)$
 u_t and v_t are generated by the processes;

$$u_{1t} = p_{11} \eta_{1t} \tag{23}$$

$$v_{1t} = p_{21} \eta_{1t} + p_{22} \eta_{2t} + p_{23} \eta_{1t-1} \tag{24}$$

The initial values are: $x_0 = 1, y_0 = 1, \alpha = 0$

$$(\beta_0, \beta_1, \phi) = B_1 = (1, 0, 0), B_2 = (0.6, 0, 0.4), B_3 = (0.6, 0.4, 0), B_4 = (0.6, 0.4, 0.4)$$

$$p_{11} = 0.95, 1.0$$

$$(p_{21}, p_{22}, p_{23}) = P_1 = (0, 1, 0), P_2 = (0.5, 0.866, 0), P_3 = (0.5, 0.866, 0.5), P_4 = (0.5, 0.9, 0.9), \\ P_5 = (0.9, 0.5, 0.5), P_6 = (0.9, 0.9, 0.9)$$

η_{1t} and η_{2t} are independently and identically distributed standard normal variables.

The various combinations of P_{2i} allow for correlation between the u_t and v_t and hence the endogeneity of X_t and for some serial correlation in v_t .

5 Results

Tables 1 and 2 present results of the persistence of the estimation of the short-run model at $P_{11} = 0.95$ and 1.0 . We observed the following:

- i. The short-run model yielded an explosive behaviour when $\phi = 0$, it seems that the adjustment of Y is too large and overshoots the equilibrium, creating a divergent pattern.
- ii. When $P_{22}, P_{23} = 0.9$ and $\phi = 0.4$ we observed that the adjustment of Y is too large and overshoots the equilibrium, creating a divergent pattern.
- iii. When $\phi \neq 0$, the adjustment of Y is gradual and convergent, and long-run will eventually approaches the equilibrium point.
- iv. The influence of the lag of the independent variable on the current value of the dependent variable continues to decrease across all β_0, β_1, ϕ .
- v. For all β_0, β_1, ϕ , the influence of the current independent variable on the current dependent variable increases for a while and then decreases.
- vi. For all the adjustment value (λ) less than -1 , it shows that the adjustment of the dependent variable Y is too large and overshoots the equilibrium, creating a divergent pattern. The error correction model fails to capture the deviation from equilibrium appropriately, thereby resulting in overshooting of the model.

Tables 3 and 4 present results of the stabilizing error correction model, we observed the following:

- i. The adjustment term was within $-1 < \lambda < 0$ which implies that there was a gradual and convergent and the long run will eventually approach the equilibrium.
- ii. The values of the coefficient of the stabilizing model was observed to be very high which implies that stabilising model is exerting a stabilizing effect on the error correction mechanism in our model, thereby reducing the overshooting in the error correction model.
- iii. The coefficient of the stabilizing model was observed to be consistently positive which implies that it contributes to a smoother and more stable response to deviations from the long run equilibrium.
- iv. The standard error of the coefficient of the adjustment term in the stabilizing error correction mechanism was observed to be smaller compared to the error correction model.
- v. There is an increase in the effect of the independent variable (x_t) on the dependent variable (y_t) in the stabilizing error correction mechanism with a smaller standard error as compared to the error correction model.

Table 1. Persistence at P11=1.0 when the sample size is 50

		N=50, P11= 1.0											
		$(\beta_0, \beta_1, \phi) = (1, 0, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0, 0.4)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0.4)$		
ESTIMATOR		COEFF	ST.ERR	P-VALUE	COEFF	ST.ERR	P-VALUE	COEFF	ST.ERR	P-VALUE	COEFF	ST.ERR	P-VALUE
P21,P22,P23 (0,1,0)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	-	-	-	-	-	-	0.7269	0.1650	0.000	0.7206	0.1631	0.000
	D(X _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	λ	-1.106	1.1244	0.000	-0.737	0.1092	0.000	-1.123	0.1429	0.000	-0.744	0.1106	0.000
P21,P22,P23 (0.5, 0.866, 0)	D(Y _{t-1})	0.1203	0.1185	0.316	-	-	-	0.1239	0.117	0.299	-	-	-
	D(X _t)	1.0452	0.1235	0.000	0.6463	0.1231	0.000	0.6409	0.1233	0.000	0.6471	0.1239	0.000
	D(X _{t-1})	0.5088	0.1684	0.004	0.5594	0.1314	0.000	0.4377	0.1860	0.023	0.4158	0.1943	0.038
	λ	-1.2634	0.1767	0.000	-0.767	0.1267	0.000	-1.295	0.1973	0.000	-0.802	0.1439	0.000
P21,P22,P23 (0.5,0.866, 0.5)	D(Y _{t-1})	-	-	-	-	-	-	0.2569	0.0759	0.002	-	-	-
	D(X _t)	1.3223	0.1039	0.000	0.9412	0.1017	0.000	0.9503	0.1019	0.000	0.9351	0.0985	0.000
	D(X _{t-1})	0.5106	0.1194	0.000	0.3209	0.1012	0.003	-	-	-	-	-	-
	λ	-1.3908	0.1267	0.000	-1.055	0.1353	0.000	-1.729	0.1679	0.000	-1.069	0.0983	0.000
P21,P22,P23 (0.5, 0.9, 0.9)	D(Y _{t-1})	-	-	-	0.0781	0.0499	0.125	0.1571	0.0615	0.012	-	-	-
	D(X _t)	1.4107	0.0771	0.000	-	-	-	1.0094	0.0757	0.000	1.0039	0.0760	0.000
	D(X _{t-1})	0.3542	0.1071	0.002	-	-	-	-	-	-	-0.191	0.1182	0.113
	λ	-1.4379	0.1407	0.000	-1.040	0.0548	0.000	-1.647	0.1785	0.000	-1.149	0.1500	0.000
P21,P22,P23 (0.9, 0.5, 0.5)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	1.0524	0.0577	0.000	0.6804	0.0576	0.000	0.6524	0.0577	0.000	0.6578	0.058	0.000
	D(X _{t-1})	0.8466	0.0634	0.000	0.5915	0.0571	0.000	0.6504	0.0563	0.000	0.2389	0.1367	0.088
	λ	-1.4907	0.0688	0.000	-1.164	0.0799	0.000	-1.491	0.0689	0.000	-1.163	0.1288	0.000
P21,P22,P23 (0.9, 0.9, 0.9)	D(Y _{t-1})	-	-	-	-	-	-	0.2276	0.0459	0.000	-	-	-
	D(X _t)	1.1762	0.0625	0.000	0.7918	0.0604	0.000	0.7844	0.0614	0.000	0.7754	0.0613	0.000
	D(X _{t-1})	0.5146	0.0771	0.000	0.2304	0.0583	0.000	-	-	-	-0.156	0.1227	0.211
	λ	-1.5627	0.1123	0.000	-1.272	0.1277	0.000	-1.894	0.1342	0.000	-1.241	0.1338	0.000

Table 2. Persistence at P11=0.95 when the sample size is 50

		N=50, P11= 0.95											
		$(\beta_0, \beta_1, \phi) = (1, 0, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0, 0.4)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0.4)$		
ESTIMATOR		COEFF	ST.ERR	P-VALUE	COEFF	ST.ERR	P-VALUE	COEFF	ST.ERR	P-VALUE	COEFF	ST.ERR	P-VALUE
P21,P22,P23 (0,1,0)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _i)	-	-	-	-	-	-	0.7678	0.1536	0.000	0.7512	0.1523	0.000
	D(X _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	λ	-1.088	0.1077	0.000	-0.713	0.1013	0.000	-1.124	0.1408	0.000	-0.723	0.1032	0.000
P21,P22,P23 (0.5, 0.866, 0)	D(Y _{t-1})	0.1416	0.1203	0.246	-	-	-	0.2779	0.0781	0.000	-	-	-
	D(X _i)	1.0553	0.1139	0.0000	0.6549	0.1137	0.000	0.7002	0.1150	0.000	0.7022	0.1123	0.000
	D(X _{t-1})	0.3986	0.1674	0.023	0.4584	0.1279	0.000	-	-	-	-	-	-
	λ	-1.3075	0.1829	0.000	-0.800	0.1326	0.000	-1.593	0.1323	0.000	-1.0019	0.0816	0.000
P21,P22,P23 (0.5, 0.866, 0.5)	D(Y _{t-1})	0.1416	0.1109	0.209	-	-	-	0.2209	0.0683	0.002	-	-	-
	D(X _i)	1.3189	0.0952	0.000	0.9301	0.0932	0.000	0.9248	0.0932	0.000	0.9182	0.0958	0.000
	D(X _{t-1})	0.3017	0.153	0.055	0.2361	0.0937	0.016	-	-	-	-0.0605	0.1549	0.699
	λ	-1.5663	0.1811	0.000	-1.090	0.1387	0.000	-1.705	0.1573	0.000	-1.0907	0.1452	0.000
P21,P22,P23 (0.5, 0.9, 0.9)	D(Y _{t-1})	0.2150	0.0705	0.004	-	-	-	0.1209	0.0541	0.031	-	-	-
	D(X _i)	1.4328	0.0679	0.000	-	-	-	0.9888	0.0696	0.000	0.9861	0.0695	0.000
	D(X _{t-1})	-	-	-	-	-	-	-	-	-	-0.2515	0.1206	0.043
	λ	-1.609	0.1847	0.000	-0.981	0.0486	0.000	-1.615	0.1677	0.000	-1.1542	0.1452	0.000
P21,P22,P23 (0.9, 0.5, 0.5)	D(Y _{t-1})	0.0827	0.0624	0.192	-	-	-	0.1060	0.0809	0.197	-	-	-
	D(X _i)	1.0623	0.0533	0.000	0.6814	0.0524	0.000	0.6604	0.0532	0.000	0.6671	0.0533	0.000
	D(X _{t-1})	0.6665	0.0879	0.000	0.4753	0.0536	0.000	0.4209	0.1162	0.001	0.15145	0.1358	0.271
	λ	-1.5878	0.0973	0.000	-1.209	0.0832	0.000	-1.656	0.1396	0.000	-1.1522	0.1231	0.000
P21,P22,P23 (0.9, 0.9, 0.9)	D(Y _{t-1})	0.1217	0.0871	0.170	-	-	-	0.1811	0.0409	0.000	-	-	-
	D(X _i)	1.1769	0.0570	0.000	0.7855	0.0553	0.000	0.7737	0.0564	0.000	0.7694	0.0565	0.000
	D(X _{t-1})	0.3302	0.1116	0.005	0.1483	0.0556	0.0108	-	-	-	-0.2064	0.1228	0.100
	λ	-1.7079	0.1494	0.000	-1.311	0.1315	0.000	-1.844	0.1251	0.000	-1.2239	0.1287	0.000

Table 3. Persistence at P11=1.0 when the sample size is 50

		N=50, P11= 1.0											
		$(\beta_0, \beta_1, \phi) = (1, 0, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0, 0.4)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0.4)$		
	ESTIMATOR	COEFF	ST.ERR	PVALUE	COEFF	ST.ERR	PVALUE	COEFF	ST.ERR	PVALUE	COEFF	ST.ERR	PVALUE
P21,P22,P23 (0,1,0)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	1.0804	0.0251	0.000				0.7134	0.0240	0.000			
	D(X _{t-1})	-	-	-				0.3019	0.0239	0.000			
	λ	-0.124	0.0313	0.000				-0.136	0.030	0.000			
	Θ	0.9814	0.0219	0.000				0.9861	0.021	0.000			
P21,P22,P23 (0.5, 0.866, 0)	D(Y _{t-1})	-	-	-				-0.036	0.0182	0.058			
	D(X _t)	1.0581	0.0219	0.000				0.6715	0.0219	0.000			
	D(X _{t-1})	0.6493	0.0214	0.000				1.0651	0.0237	0.000			
	λ	-0.117	0.0373	0.003				-0.118	0.0414	0.007			
	Θ	1.0234	0.0265	0.000				1.0169	0.0266	0.000			
P21,P22,P23 (0.5, 0.866, 0.5)	D(Y _{t-1})	-	-	-	-	-	-	-0.293	0.0547	0.000	-	-	-
	D(X _t)	1.3418	0.070	0.000	1.0652	0.0768	0.000	1.0600	0.064	0.000	0.9493	0.011	0.000
	D(X _{t-1})	0.0322	0.0717	0.666	0.2249	0.0755	0.005	0.5457	0.0916	0.000	0.7942	0.011	0.000
	λ	-0.481	0.1425	0.002	-0.399	0.1503	0.011	-0.181	0.1336	0.182	-0.058	0.023	0.015
	Θ	0.906	0.0980	0.000	1.063	0.1083	0.000	0.986	0.0913	0.000	1.008	0.016	0.000
P21,P22,P23 (0.5, 0.9, 0.9)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	1.454	0.0538	0.000	1.099	0.0077	0.00	1.112	0.0509	0.000	1.023	0.014	0.000
	D(X _{t-1})	-0.332	0.0536	0.000	-	-	-	-0.002	0.0486	0.9614	0.5764	0.013	0.000
	λ	-0.486	0.1513	0.003	0.0522	0.0244	0.038	-0.431	0.1443	0.005	-0.154	0.0379	0.000
	Θ	0.924	0.1019	0.000	1.007	0.0158	0.000	0.996	0.0998	0.000	0.983	0.0268	0.000
P21,P22,P23 (0.9, 0.5, 0.5)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	1.117	0.0631	0.000	0.711	0.020	0.000	0.738	0.0949	0.000	0.680	0.0130	0.000
	D(X _{t-1})	0.682	0.0663	0.000	0.956	0.0214	0.000	1.013	0.0977	0.000	1.411	0.0138	0.000
	λ	-0.615	0.2287	0.010	-0.101	0.0705	0.160	-0.690	0.3441	0.051	-0.173	0.0473	0.001
	Θ	0.761	0.1592	0.000	1.069	0.0525	0.000	0.768	0.2417	0.003	0.985	0.0343	0.000
P21,P22,P23 (0.9, 0.9, 0.9)	D(Y _{t-1})	-	-	-	-	-	-	-0.532	0.0615	0.000	-	-	-
	D(X _t)	1.258	0.0684	0.000	0.838	0.0243	0.000	0.878	0.0539	0.000	0.811	0.0208	0.000
	D(X _{t-1})	0.090	0.0710	0.211	0.455	0.0254	0.000	0.894	0.0913	0.000	0.918	0.0219	0.000
	λ	-0.689	0.2331	0.005	-0.232	0.0814	0.007	-0.101	0.1825	0.583	-0.245	0.0737	0.002
	Θ	0.774	0.1599	0.000	1.020	0.0591	0.000	0.992	0.1287	0.000	0.976	0.0516	0.000

Table 4. Persistence at P11=0.95 when the sample size is 50

		N=50, P11= 0.95											
		$(\beta_0, \beta_1, \phi) = (1, 0, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0, 0.4)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0)$			$(\beta_0, \beta_1, \phi) = (0.6, 0.4, 0.4)$		
	ESTIMATOR	COEFF	ST.ERR	PVALUE	COEFF	ST.ERR	PVALUE	COEFF	ST.ERR	PVALUE	COEFF	ST.ERR	PVALUE
P21,P22,P23 (0,1,0)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	1.171	0.0202	0.000	-	-	-	0.762	0.0230	0.000	-	-	-
	D(X _{t-1})	-	-	-	-	-	-	0.368	0.0228	0.000	-	-	-
	λ	-0.101	0.0277	0.0007	-	-	-	-0.138	0.0312	0.0001	-	-	-
	Θ	0.987	0.0191	0.000	-	-	-	0.986	0.0218	0.000	-	-	-
P21,P22,P23 (0.5, 0.866, 0)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	1.071	0.0237	0.000	-	-	-	0.757	0.069	0.000	0.703	0.0004	0.000
	D(X _{t-1})	0.614	0.0231	0.000	-	-	-	0.864	0.0068	0.000	1.138	0.0004	0.000
	λ	-0.142	0.0436	0.0023	-	-	-	-0.255	0.115	0.0324	-0.002	0.0007	0.006
	Θ	1.022	0.0312	0.000	-	-	-	1.083	0.0883	0.000	0.999	0.0005	0.000
P21,P22,P23 (0.5, 0.866, 0.5)	D(Y _{t-1})	-	-	-	-	-	-	-	-	-	-	-	-
	D(X _t)	1.369	0.0568	0.000	0.942	0.0102	0.000	1.036	0.0764	0.000	0.934	0.0103	0.000
	D(X _{t-1})	0.044	0.0582	0.451	0.434	0.0105	0.000	0.286	0.0748	0.0004	0.881	0.0102	0.000
	λ	-0.402	0.125	0.003	-0.09	0.0222	0.000	-0.428	0.1641	0.013	-0.103	0.0222	0.000
	Θ	1.007	0.0879	0.000	0.996	0.0158	0.000	1.042	0.1172	0.000	0.985	0.0158	0.000
P21,P22,P23 (0.5, 0.9, 0.9)	D(Y _{t-1})	-	-	-	-	-	-	-0.054	0.0319	0.0965	-	-	-
	D(X _t)	1.527	0.0425	0.000	-	-	-	1.119	0.0466	0.00	1.005	0.0132	0.000
	D(X _{t-1})	-0.406	0.0409	0.000	-	-	-	-	-	-	0.603	0.0131	0.000
	λ	-0.326	0.134	0.019	-	-	-	-0.377	0.1536	0.0182	-0.159	0.0405	0.000
	Θ	1.014	0.0902	0.000	-	-	-	0.983	0.1025	0.000	0.982	0.0284	0.000
P21,P22,P23 (0.9, 0.5, 0.5)	D(Y _{t-1})	-0.445	0.0096	0.000	-	-	-	-0.443	0.0097	0.000	-	-	-
	D(X _t)	1.076	0.0114	0.000	0.718	0.0244	0.000	0.681	0.0114	0.000	0.677	0.0116	0.000
	D(X _{t-1})	1.196	0.0175	0.000	0.916	0.0249	0.000	1.411	0.0149	0.000	1.381	0.0122	0.000
	λ	-0.039	0.0449	0.394	-0.16	0.0903	0.082	-0.095	0.0444	0.039	-0.163	0.0467	0.001
	Θ	1.016	0.0031	0.000	1.054	0.0677	0.000	1.006	0.0316	0.000	0.986	0.0332	0.000
P21,P22,P23 (0.9, 0.9, 0.9)	D(Y _{t-1})	-	-	-	-	-	-	-0.574	0.0509	0.000	-	-	-
	D(X _t)	1.327	0.0523	0.000	0.837	0.0256	0.000	0.844	0.0426	0.000	0.801	0.0184	0.000
	D(X _{t-1})	-0.639	0.1009	0.000	0.441	0.0267	0.000	0.972	0.0732	0.000	0.918	0.0192	0.000
	λ	-0.466	0.2155	0.0362	-0.29	0.0936	0.0032	-0.019	0.1583	0.9054	-0.225	0.0715	0.0031
	Θ	1.029	0.1509	0.000	0.995	0.0684	0.000	1.036	0.1096	0.000	0.980	0.0493	0.000

5.1 Comparison between error correction model and adjustment for explosives

Table 5 shows the comparison between error correction model and the adjustment for explosiveness. It was discovered that the adjustment value from the error correction model was explosive but it became stabilized and ranges from -0.7 to 0.5 in the stabilizing error correction model. Furthermore, the Sum of Square of regression for the stabilizing error correction model was observed to be smaller than the Sum of square of regression of the error correction model.

Table 5. Comparison between error correction model and stabilizing error correction model

P11=1.0, N=50							
	Adjustment term (λ) of ECM	SSR of ECM	Adjustment term (λ) of SECM	θ of SECM	P-Value of λ of SECM	P-value of θ of SECM	SSR of θ of SECM
B1P1	-1.10655	60.6888	-0.12440	0.98137	0.000	0.000	1.30663
B1P2	-1.26337	35.5929	-0.11647	1.02339	0.003	0.000	1.04553
B1P3	-1.3908	31.5238	-0.48134	0.90574	0.002	0.000	12.55786
B1P4	-1.437906	24.31958	-0.48642	0.92381	0.003	0.000	10.49876
B1P5	-1.490685	10.14827	-0.615326	0.476059	0.010	0.000	10.0034
B1P6	-1.56265	19.76874	-0.68851	0.774112	0.005	0.000	19.41545
B2P3	-1.054627	30.55614	-0.05275	1.000434	0.000	0.000	0.116955
B2P4	-1.040132	24.24044	0.052218	1.00697	0.038	0.000	0.255406
B2P5	-1.16411	10.32825	-0.100937	1.06846	0.159	0.000	1.172805
B2P6	-1.271669	19.0875	-0.231787	1.02035	0.007	0.000	2.76766
B3P1	-1.12293	60.0683	-0.135836	0.986107	0.000	0.000	1.1716
B3P2	-1.295019	36.51251	-0.118448	1.016958	0.007	0.000	1.02564
B3P3	-1.72957	31.37935	-0.39903	1.062714	0.011	0.000	15.12816
B3P4	-1.64704	23.79466	-0.431338	0.99639	0.005	0.000	10.18373
B3P5	-1.490685	10.14827	-0.690228	0.76807	0.051	0.003	24.27447
B3P6	-1.8940	19.4337	-0.10106	0.99155	0.583	0.000	13.29955
B4P3	-1.068969	30.86852	-0.05828	1.00814	0.015	0.000	0.3488
B4P4	-1.149705	21.95608	-0.15423	0.983161	0.000	0.000	0.61719
B4P5	-1.16268	9.29939	-0.17327	0.98506	0.001	0.000	0.43112
B4P6	-1.240767	17.48065	-0.24517	0.97566	0.002	0.000	1.851388
P11=0.95 and N=50							
	Adjustment term (λ) of ECM	SSR of ECM	Adjustment term (λ) of SECM	θ of SECM	P-Value of λ of SECM	P-value of θ of SECM	SSR of θ of SECM
B1P1	-1.08815	54.14689	-0.10133	0.98666	0.001	0.000	0.87679
B1P2	-1.30745	32.56275	-0.14165	1.02145	0.002	0.000	1.27827
B1P3	-1.5663	27.25855	-0.40192	1.006948	0.003	0.000	8.72927
B1P4	-1.60915	21.91993	-0.32638	1.013603	0.019	0.000	7.50795
B1P5	-1.58782	8.79073	-0.0387	1.01632	0.394	0.000	0.3462
B1P6	-1.7079	16.99948	-0.46628	1.028523	0.036	0.000	16.05396
B2P3	-1.08961	27.00202	-0.09139	0.99576	0.000	0.000	0.28104
B2P5	-1.20909	8.9427	-0.16071	1.053848	0.082	0.000	1.67437
B2P6	-1.31123	16.86052	-0.29248	0.99493	0.003	0.000	3.21218
B3P1	-1.12398	53.77957	-0.137669	0.985928	0.000	0.000	1.108223
B3P2	-1.59286	34.5018	-0.2550	1.08251	0.032	0.000	10.42417
B3P3	-1.70537	27.43358	-0.4275	1.042278	0.013	0.000	15.6464
B3P4	-1.61517	21.17412	-0.37719	0.98297	0.018	0.000	9.54722
B3P5	-1.65556	8.79209	0.09494	1.006336	0.004	0.000	0.35042
B3P6	-1.84399	17.14969	-0.01894	1.03599	0.905	0.000	8.571186
B4P2	-1.001949	34.5340	-0.001997	0.99999	0.006	0.000	0.00041
B4P3	-1.09065	25.52617	-0.10264	0.98456	0.000	0.000	0.24994
B4P4	-1.15423	19.3038	-0.15957	0.98158	0.000	0.000	0.61537
B4P5	-1.15222	8.12881	-0.16291	0.98564	0.001	0.000	0.355688
B4P6	-1.22386	15.51496	-0.22498	0.980316	0.003	0.000	1.4984

6 Conclusion

In modeling the short-run relationship, we observed that most of the adjustment terms were less than -1, which implies that the adjustment term is explosive, thereby creating an over-correction, and it also implies an oscillatory convergence. To correct the problem of explosiveness in the short-run model, a stabilizing error correction mechanism was proposed. The coefficient of the stabilizing error correction model was observed to be consistently positive, with values greater than 0 in all cases, which means that the damping coefficient exerted a stabilizing effect on the error correction mechanism in the model, thereby reducing the overshooting in the error correction model. It also implies that the adjustment mechanism responds to deviations from the long-run equilibrium in a way that prevents rapid and excessive corrections. It contributes to a smoother and more stable response to deviations from the long-run equilibrium. There is an increase in the effect of the independent variable (x_t) on the dependent variable (y_t) in the stabilizing error correction mechanism with a smaller standard error as compared to the error correction model.

The root mean square error of the stabilizing Error Correction model is observed to be smaller than the adjustment model in the Error correction model. The stabilizing error correction model performs better than the Error correction model.

Competing Interests

Authors have declared that no competing interests exist.

References

- [1] Pesaran MH, Sin Y. An Autoregressive Distributed lag modeling approach to Cointegration analysis. University of Cambridge; 1997.
- [2] Kripfganz S, Schneider DC. ARDL: Estimating autoregressive distributed lag and equilibrium correction models. Tohoku University Research Center for Policy Design Discussion Paper. 2022;6.
- [3] Pesaran MH, Shin Y, Smith RJ. Bounds testing approaches to the analysis of level relationships. *Journal of Applied Econometrics*. 2001;16(3):289-326.
- [4] Adamu A, Usman HH. Does oil price and production influence economic growth? Evidence from Nigeria. *Journal of Economics and Allies Research*. 2022;7(2).
- [5] ISSN: 2536-7447.
- [6] Aronu CO, Ekwueme OG, Bilesanmi AO, Arowolo OT. The nexus between some selected macroeconomic variables on growth output in Nigeria: An autoregressive distributed lag model approach. *Economics and Business Quarterly Reviews*. 2022;5(4):256-262.
- [7] DOI: 10.31014/aior.1992.05.04.474
- [8] Ibrahim VH. Monetary policy and economic growth in Nigeria An Autoregressive Distributed Lag (ARDL) Model, *Advances in Social Sciences Research Journal*. 2019;6(3):90-100.
- [9] Charles Odinakachi Njoku, Chilaka Emmanuel Nwaimo. The impact of exchange rate on inflation in Nigeria. *Management Studies and Economic Systems*. 2019; 4(3):171-195.
- [10] Elem-Uche, Omekara CD, Okereke EW, Madu C. Vector error correction model for forecasting real output using monetary policy transmission channel variables for Nigeria. *American Journal of Mathematics and Statistics*. 2019;9(2):92-99.
- [11] DOI: 10.5923/j.ajms.20190902.04

- [12] Adeyele, N., E. Osabuohien, E. Bowale, O. Matthew, and E. Oduntan. Financial reforms and credit growth in Nigeria: Empirical insights from ARDL and ECM techniques. *International Review of Applied Economics* 2018, 32(6): 807–820.
- [13] Narayan PK, Smyth R. The residential demand for electricity in Australia: An application of the bounds testing approach to cointegration. *Energy Policy*, 2005;33(4):467-474.
- [14] Cantah WG, Asmah EE Crude oil price and growth of output: the case of Ghana. *International Journal of Economics, Commerce and Management*. 2015;3(5):470-498.
- [15] Mustafa I. Effects of monetary policy on macroeconomic performance: The case of Nigeria (Doctoral Dissertation). University of Greenwich, London; 2013
- [16] Celina UC. Monetary policy and Economic growth in Nigeria. *Journal of Policy and Development Studies*. 2014;9(1):234-247.
- [17] Enisan, A. A., and A. O. Olufisayo. Stock market development and economic growth: Evidence from seven sub-Sahara African countries. *Journal of Economics and Business*. 2009, 61(2): 162–171.
- [18] Rostin Muthalib AA, Adam P, Nur M, Saenong Z, Suriadi LO, Baso JN. The effect of crude oil prices on inflation, interest rates and economic growth in Indonesia. *International Journal of Energy Economic and Policy*. 2019;9(5):14-19.
- [19] Available:<http://doi.org/10.324doi.orhdl.handle.et/11159/5077>
- [20] Liaqat M, Ashraf A, Nisar S, Khursheed A. The impact of oil price inflation on economic growth of oil importing economics: Empirical evidence from Pakistan. *Journal of Asian Finance, Economics and Business*. 2022;9(1):167-176.
- [21] Available:<http://doi.10.13106/jafeb.2022.vol9.no1.0167>
- [22] Bahmani-Oskooee M, Narayan PK. An empirical re-examination of the long-run effects of trade on income. *Journal of Economic Development*. 2006;31(1):1-11
- [23] Belloumi M. The relationship between trade, FDI and economic growth in Tunisia: An application of the autoregressive distributed lag model. *Economic Systems*. 2014;38(2):269–287
- [24] Ozturk I, Acaravci A. The long-run and causal analysis of energy, growth, openness and financial development on carbon emissions in Turkey. *Energy Economics*. 2013;36:262-267.
- [25] Hamid S, Tartar M, Mohtashami S. Impact of determinant macro economic variables on environmental changes in Iran. *Environ Health Insights* 2023;17.
- [26] Available:<https://doi/10.1177/11786302221149855>
- [27] Saida Z, Kais S. Environmental pollution, health expenditure and economic growth in the sub-Saharan African Countries: Panel ARDL approach. *Science Direct*. 2018;41:833-840.
- [28] Mamun A, Sohaq K. The effect of economic growth, urbanization, and globalization on energy consumption. A panel data analysis of BRICS countries, *Energy Sources, Part B. Economics, Planning and Policy*. 2017;12(3):210-218
- [29] Zhigang H, Huang L. Individual new energy consumption and economic growth in China. *The North American Journal of Economics and Finance*. 2019;(54).
- [30] Available:<https://doi.org/10.1016/j.najet.2019.101010>
- [31] Box GEP, Jenkins GM. *Time series analysis: Forecasting and control*, Holden Day: San Francisco; 1970.
- [32] Engle RF, Granger CWJ. Cointegration and error correction: representation, estimation and testing, *Econometrica*. 1987;55:251-76.

- [33] Iyeli I, Iyeli Anthony I, Aham KU. Cointegration and econometric analysis of non-stationary data in Nigeria: An empirical evidence. *Journal of Contemporary Research*. 2011;8(1):194-213.
- [34] ISSN: 1813-2227
- [35] Nkoro E, Uko AK. Autoregressive Distributed Lag (ARDL) cointegration technique: Application and interpretation. *Journal of Statistical and Econometric Methods*. 2016;5(4).
- [36] Muritala T, Taiwo A, Olowookere D. Crude oil price, Stock price and some selected macroeconomic indicators: Implication on the growth of Nigeria Economy. *Research Journal of Financial and Accounting*. 2012;3(2):42-49.
- [37] ISSN 2222-2847
- [38] Abdulaziz HA, Basmah SA. Monetary volatility and the dynamics of economic growth: An empirical analysis of an oil-based economy, 1970-2018. *OPEC Energy Review*. 2022;47(1):3-18.
- [39] Wasanthi M. Money supply, inflation and economic growth of Sri Lanka Co-integration and causality analysis. *Journal of Money and Business*. 2023;3(2):227-236
- [40] Stock JH. Asymptotic properties of least square estimators of cointegrating vectors. *Econometrics*. 1987;55:1035-1056.
- [41] Perasan MH, Smith RJ. Structural analysis of cointegration VARs, *Journal of Economic Surveys*. 1998;12(5):471-505.

© 2024 Abifarin et al.; This is an Open Access article distributed under the terms of the Creative Commons Attribution License (<http://creativecommons.org/licenses/by/4.0>), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Peer-review history:

The peer review history for this paper can be accessed here (Please copy paste the total link in your browser address bar)

<https://www.sdiarticle5.com/review-history/112368>